

THE NUMBER OF 2-PROTECTED NODES IN TRIES AND PATRICIA TRIES

Guan-Ru Yu

Department of Applied Mathematics
National Chiao Tung University

August 2nd, 2014

Three main classes of digital trees

- Digital search trees (DSTs)
- Tries
- PATRICIA tries

Digital search trees (DSTs)

$$R_1 = 000001 \dots$$

$$R_2 = 000110 \dots$$

$$R_3 = 110111 \dots$$

$$R_4 = 011011 \dots$$

$$R_5 = 100001 \dots$$

$$R_6 = 111110 \dots$$

Digital search trees (DSTs)

$$R_1 = 000001 \dots$$

$$R_2 = 000110 \dots$$

$$R_3 = 110111 \dots$$

$$R_4 = 011011 \dots$$

$$R_5 = 100001 \dots$$

$$R_6 = 111110 \dots$$

R_1

Digital search trees (DSTs)

$$R_1 = 000001 \dots$$

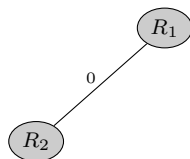
$$R_2 = 000110 \dots$$

$$R_3 = 110111 \dots$$

$$R_4 = 011011 \dots$$

$$R_5 = 100001 \dots$$

$$R_6 = 111110 \dots$$



Digital search trees (DSTs)

$$R_1 = 000001 \dots$$

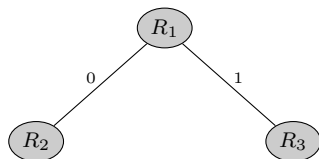
$$R_2 = 000110 \dots$$

$$R_3 = 110111 \dots$$

$$R_4 = 011011 \dots$$

$$R_5 = 100001 \dots$$

$$R_6 = 111110 \dots$$



Digital search trees (DSTs)

$$R_1 = 000001 \dots$$

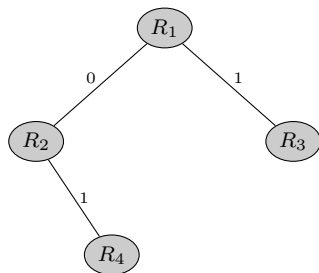
$$R_2 = 000110 \dots$$

$$R_3 = 110111 \dots$$

$$R_4 = 011011 \dots$$

$$R_5 = 100001 \dots$$

$$R_6 = 111110 \dots$$



Digital search trees (DSTs)

$$R_1 = 000001 \dots$$

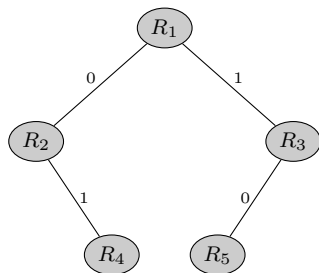
$$R_2 = 000110 \dots$$

$$R_3 = 110111 \dots$$

$$R_4 = 011011 \dots$$

$$R_5 = 100001 \dots$$

$$R_6 = 111110 \dots$$



Digital search trees (DSTs)

$$R_1 = 000001 \dots$$

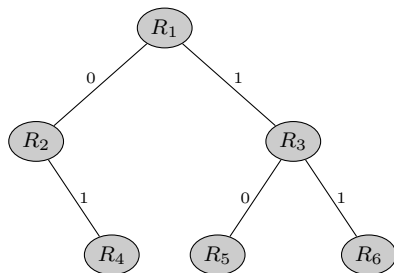
$$R_2 = 000110 \dots$$

$$R_3 = 110111 \dots$$

$$R_4 = 011011 \dots$$

$$R_5 = 100001 \dots$$

$$R_6 = 111110 \dots$$



Tries

$R_1 = 000001 \dots$

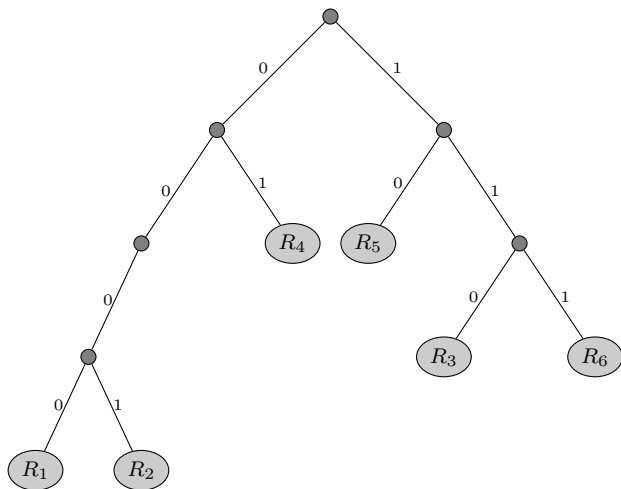
$R_2 = 000110 \dots$

$R_3 = 110111 \dots$

$R_4 = 011011 \dots$

$R_5 = 100001 \dots$

$R_6 = 111110 \dots$



PATRICIA tries

$R_1 = 000001 \dots$

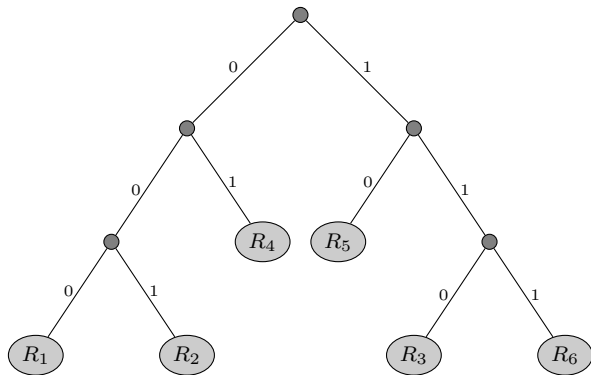
$R_2 = 000110 \dots$

$R_3 = 110111 \dots$

$R_4 = 011011 \dots$

$R_5 = 100001 \dots$

$R_6 = 111110 \dots$

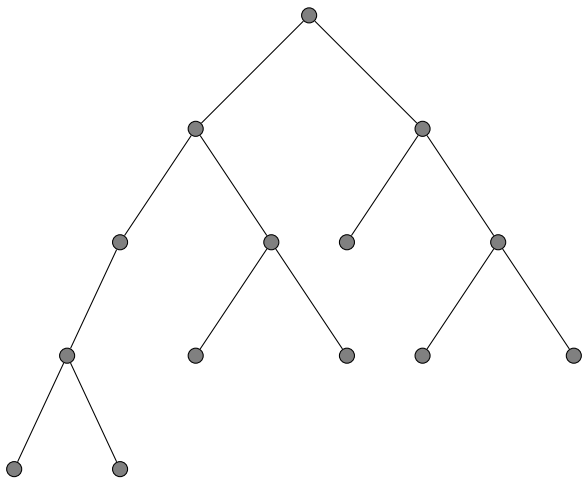


Random model

- (In probability theory) A sequence of RVs is independent and identically distributed (iid): same probability distribution and all are mutually independent.
- X_1, X_2, \dots is a random string: X_1, X_2, \dots is an iid sequence of RVs with $P(X_n = 0) = p$ and $P(X_n = 1) = q := 1 - p$.
- A digital tree is a random digital tree of size n : constructed from n infinite $\{0, 1\}$ -random strings tries.

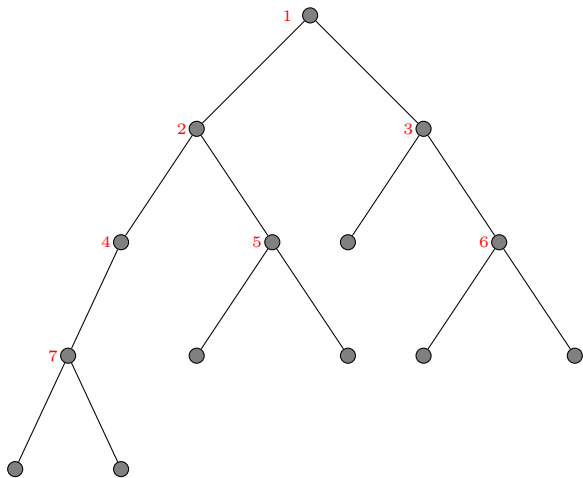
Size

Size of digital trees: number of internal nodes.



Size

Size of digital trees: number of internal nodes.



Size of tries

- Mean: Knuth (1973).

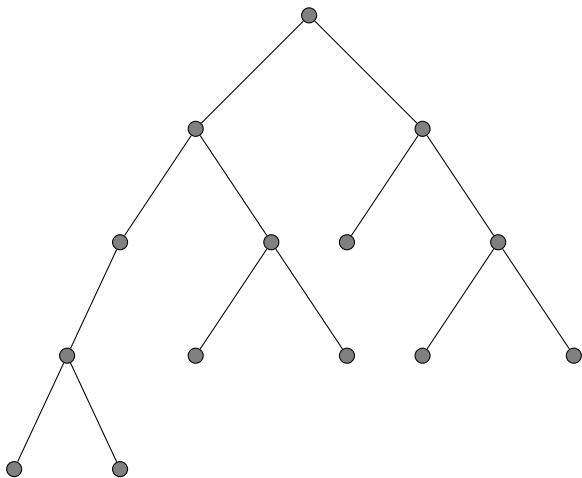
Size of tries

- Mean: Knuth (1973).
- Variance: Two groups independently.
- 1. P. Kirschenhofer and H. Prodinger: symmetric case ($p = q = 1/2$) with explicit expressions for involved constants and periodic functions (1991).
- 2. P. Jacquet and M. Regnier: general case (symmetric and asymmetric case) but without explicit expressions for involved constants and periodic functions (1988)

partial results on explicit expressions of involved constants and periodic functions (1989).

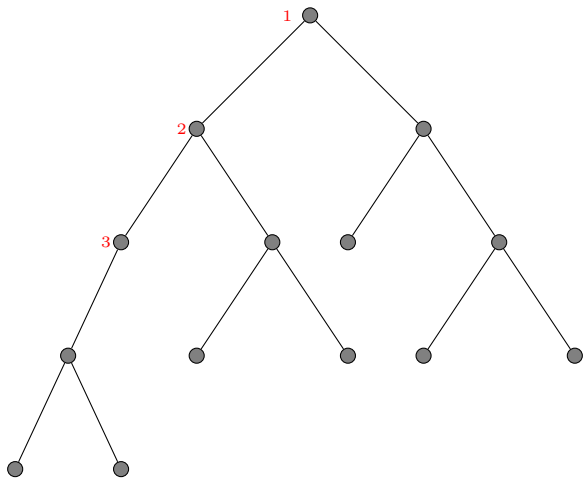
2-protected nodes

2-protected nodes: nodes that have distance at least 2 from leaves.



2-protected nodes

2-protected nodes: nodes that have distance at least 2 from leaves.



2-protected nodes in tries

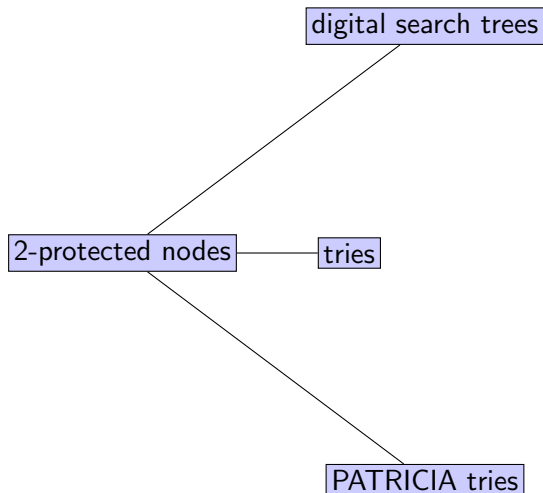
- Mean: J. Gaither, Y. Homma, M. Sellke and M. D. Ward (2012).

2-protected nodes in tries

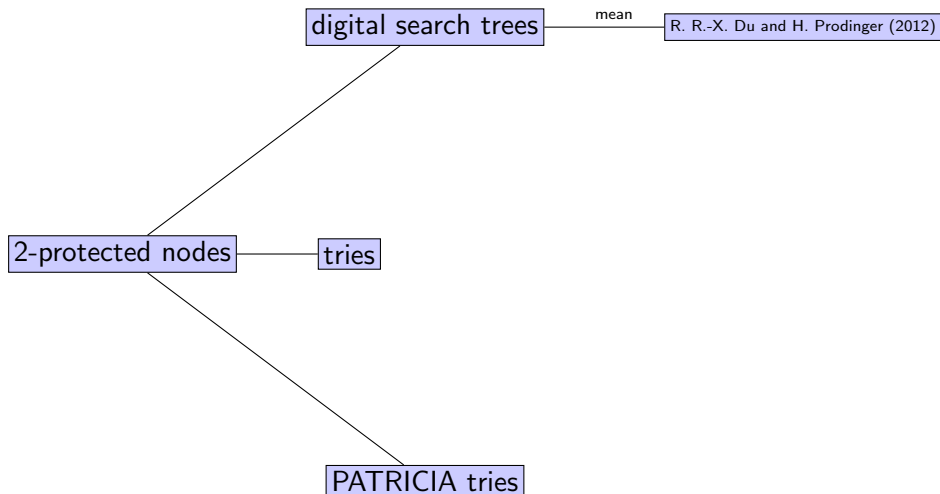
- Mean: J. Gaither, Y. Homma, M. Sellke and M. D. Ward (2012).
- Variance: J. Gaither and M. D. Ward (2013).
- Some mistakes:
 1. wrong error term.
 2. forgot to pull out the average value.

Number of 2-protected nodes in digital tree

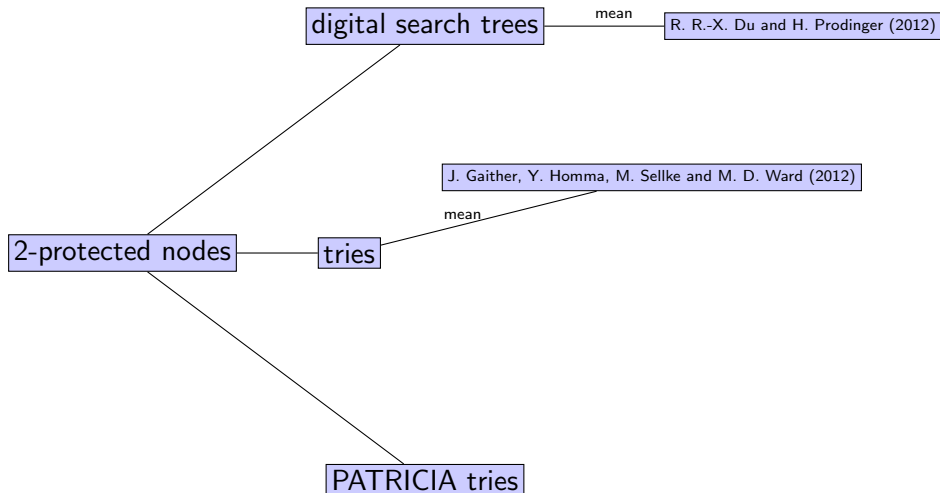
Number of 2-protected nodes in digital tree



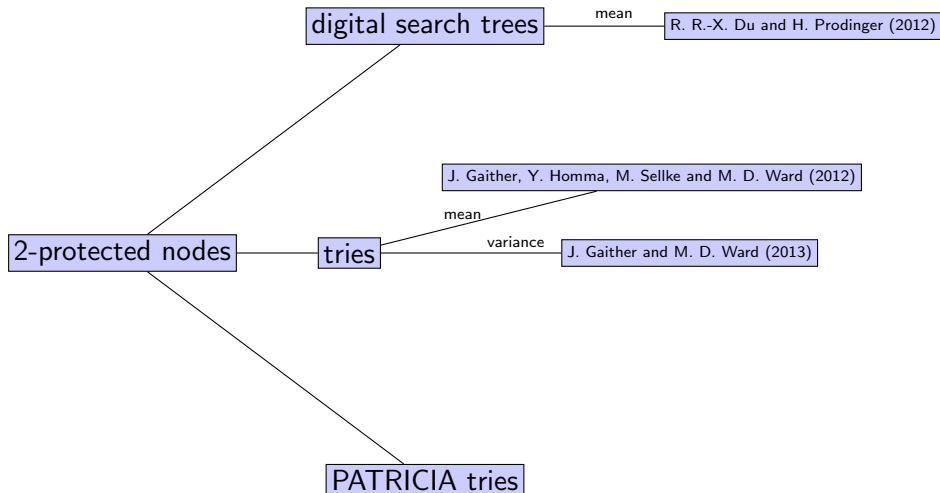
Number of 2-protected nodes in digital tree



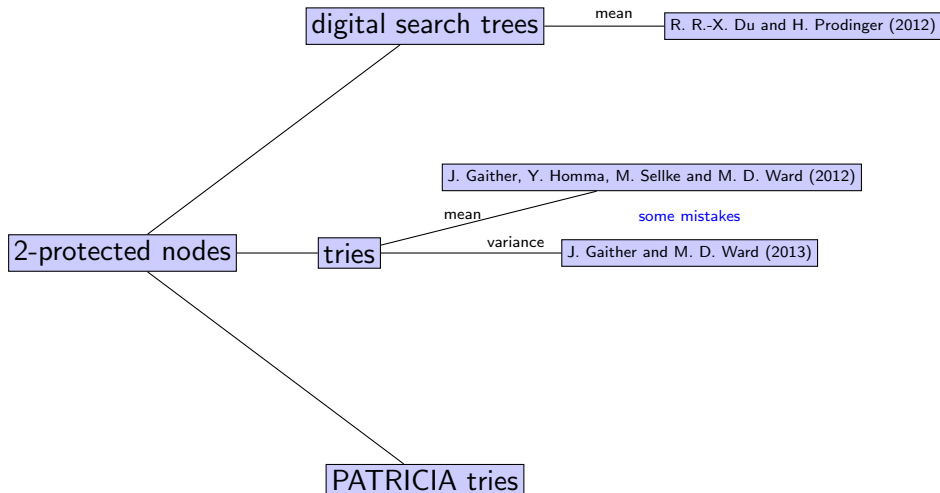
Number of 2-protected nodes in digital tree



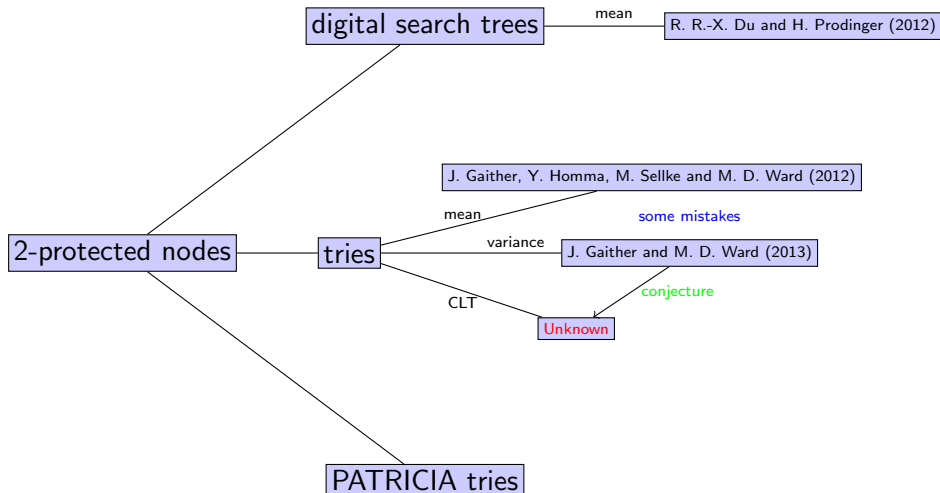
Number of 2-protected nodes in digital tree



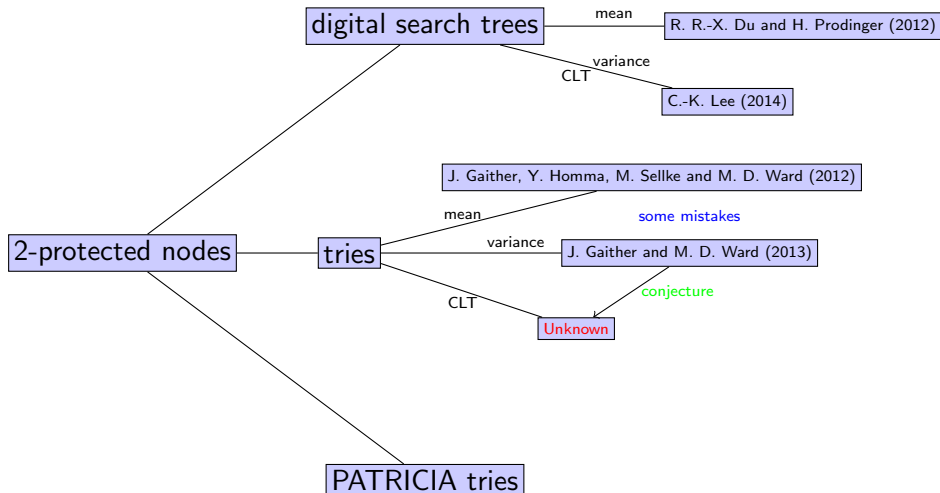
Number of 2-protected nodes in digital tree



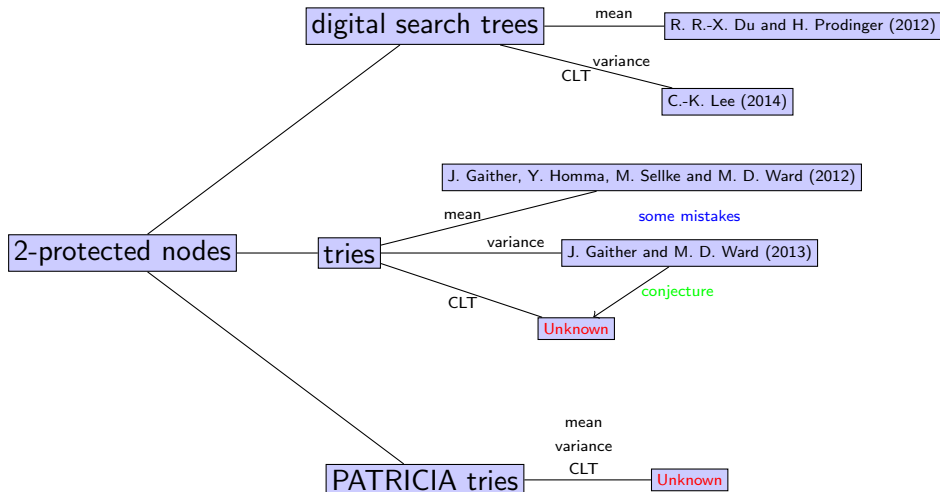
Number of 2-protected nodes in digital tree



Number of 2-protected nodes in digital tree



Number of 2-protected nodes in digital tree



Main goal of thesis

- Re-derive (and correcte) the results of J. Gaither, Y. Homma, M. Sellke and M. D. Ward by using the method of M. Fuchs, H.-K. Hwang, and V. Zacharovas.

Main goal of thesis

- Re-derive (and correct) the results of J. Gaither, Y. Homma, M. Sellke and M. D. Ward by using the method of M. Fuchs, H.-K. Hwang, and V. Zacharovas.
- Prove the conjectured central limit theorem.

Main goal of thesis

- Re-derive (and correct) the results of J. Gaither, Y. Homma, M. Sellke and M. D. Ward by using the method of M. Fuchs, H.-K. Hwang, and V. Zacharovas.
- Prove the conjectured central limit theorem.
- Derive a bivariate central limit theorem of the number of internal nodes and the number of 2-protected nodes in random tries. This result contains the central limit theorem for PATRICIA tries.

Main goal of thesis

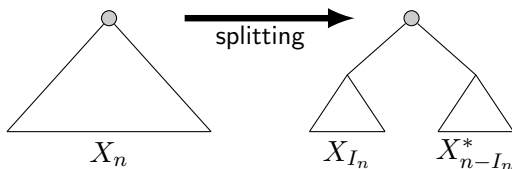
- Re-derive (and correct) the results of J. Gaither, Y. Homma, M. Sellke and M. D. Ward by using the method of M. Fuchs, H.-K. Hwang, and V. Zacharovas.
- Prove the conjectured central limit theorem.
- Derive a bivariate central limit theorem of the number of internal nodes and the number of 2-protected nodes in random tries. This result contains the central limit theorem for PATRICIA tries.
- Derive asymptotic expansions for the number of 2-protected nodes in PATRICIA tries.

Additive shape parameters

- A general framework for obtaining asymptotic expansions of mean and variance of **additive shape parameter** in random tries with explicit expressions for periodic functions.

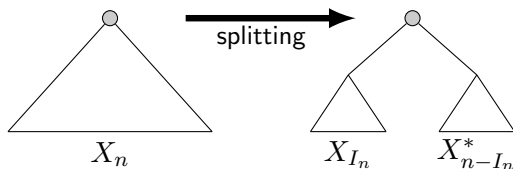
Additive shape parameters

- A general framework for obtaining asymptotic expansions of mean and variance of **additive shape parameter** in random tries with explicit expressions for periodic functions.



Additive shape parameters

- A general framework for obtaining asymptotic expansions of mean and variance of **additive shape parameter** in random tries with explicit expressions for periodic functions.



- By splitting the trie,

$$X_n \stackrel{d}{=} X_{I_n} + X_{n-I_n}^* + T_n.$$

Main steps of the analysis

- The first step is to take moments on both sides of the distributional recurrence for X_n . This gives

$$a_n = \sum_{0 \leq k \leq n} \pi_{n,k} (a_k + a_{n-k}) + b_n$$

for all moments, where b_n is a function of moments of lower order.

Main steps of the analysis

- The first step is to take moments on both sides of the distributional recurrence for X_n . This gives

$$a_n = \sum_{0 \leq k \leq n} \pi_{n,k} (a_k + a_{n-k}) + b_n$$

for all moments, where b_n is a function of moments of lower order.

- The second step is poissonization. Let

$$\tilde{f}(z) := e^{-z} \sum_n a_n \frac{z^n}{n!}, \quad \tilde{g}(z) := e^{-z} \sum_n b_n \frac{z^n}{n!}.$$

Main steps of the analysis

- The first step is to take moments on both sides of the distributional recurrence for X_n . This gives

$$a_n = \sum_{0 \leq k \leq n} \pi_{n,k} (a_k + a_{n-k}) + b_n$$

for all moments, where b_n is a function of moments of lower order.

- The second step is poissonization. Let

$$\tilde{f}(z) := e^{-z} \sum_n a_n \frac{z^n}{n!}, \quad \tilde{g}(z) := e^{-z} \sum_n b_n \frac{z^n}{n!}.$$

- We have

$$\tilde{f}(z) = \tilde{f}(pz) + \tilde{f}(qz) + \tilde{g}(z).$$

Main steps of the analysis

- The third step is doing Mellin transform on both sides. Therefore

$$\mathcal{M}[\tilde{f}(z); s] := \int_0^{\infty} \tilde{f}(z) z^{s-1} dz.$$

Main steps of the analysis

- The third step is doing Mellin transform on both sides. Therefore

$$\mathcal{M}[\tilde{f}(z); s] := \int_0^{\infty} \tilde{f}(z) z^{s-1} dz.$$

- We have

$$\mathcal{M}[\tilde{f}(z); s] = \frac{\mathcal{M}[\tilde{g}(z); s]}{1 - p^{-s} - q^{-s}}$$

Main steps of the analysis

- The third step is doing Mellin transform on both sides. Therefore

$$\mathcal{M}[\tilde{f}(z); s] := \int_0^{\infty} \tilde{f}(z) z^{s-1} dz.$$

- We have

$$\mathcal{M}[\tilde{f}(z); s] = \frac{\mathcal{M}[\tilde{g}(z); s]}{1 - p^{-s} - q^{-s}}$$

- The fourth step is using inverse Mellin transform

$$\tilde{f}(z) = \frac{1}{2\pi i} \int_{\uparrow} \frac{\mathcal{M}[\tilde{g}(z); s] z^{-s}}{1 - p^{-s} - q^{-s}} ds$$

and applying the converse mapping theorem.

Main steps of the analysis

- The last step is depoissonization which is based on the Poisson heuristic

$$\tilde{f}(n) = \mathbb{E}(a_N) \sim a_n.$$

Main steps of the analysis

- The last step is depoissonization which is based on the Poisson heuristic

$$\tilde{f}(n) = \mathbb{E}(a_N) \sim a_n.$$

- To make this step precise, we use Cauchy's integral formula

$$a_n = \frac{n!}{2\pi i} \oint_{|z|=r} z^{-n-1} e^z \tilde{f}(z) dz$$

and the saddle-point method.

Main steps of the analysis

- The last step is depoissonization which is based on the Poisson heuristic

$$\tilde{f}(n) = \mathbb{E}(a_N) \sim a_n.$$

- To make this step precise, we use Cauchy's integral formula

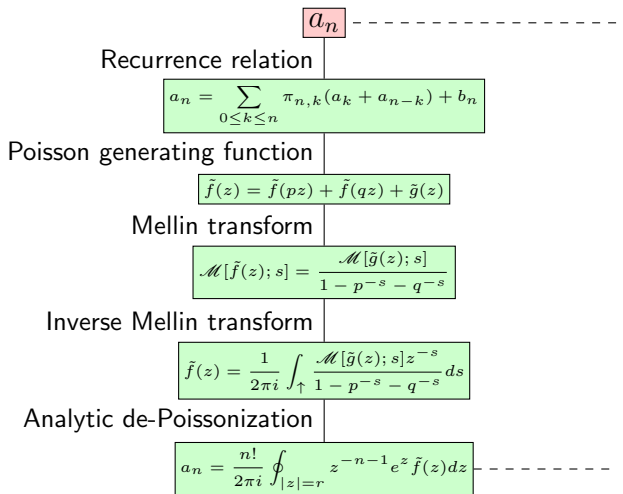
$$a_n = \frac{n!}{2\pi i} \oint_{|z|=r} z^{-n-1} e^z \tilde{f}(z) dz$$

and the saddle-point method.

- All the above steps can be merged via

$$a_n = \frac{n!}{2\pi i} \int_{\uparrow} \frac{\mathcal{M}[\tilde{g}(z); s]}{(1 - p^{-s} - q^{-s})\Gamma(n + 1 - s)} ds.$$

Poisson-Mellin-Newton cycle



Variance

- The definition of the variance is

$$\mathbb{V}(X_n) = \mathbb{E}(X_n^2) - (\mathbb{E}(X_n))^2.$$

Variance

- The definition of the variance is

$$\mathbb{V}(X_n) = \mathbb{E}(X_n^2) - (\mathbb{E}(X_n))^2.$$

- P. Jacquet and M. Regnier used the following

$$\tilde{W}(z) := \tilde{f}_2(z) - \tilde{f}_1(z)^2$$

Variance

- The definition of the variance is

$$\mathbb{V}(X_n) = \mathbb{E}(X_n^2) - (\mathbb{E}(X_n))^2.$$

- P. Jacquet and M. Regnier used the following

$$\tilde{W}(z) := \tilde{f}_2(z) - \tilde{f}_1(z)^2$$

- M. Fuchs, H.-K. Hwang and V. Zacharovas pointed out that a better choice is

$$\tilde{V}(z) := \tilde{f}_2(z) - \tilde{f}_1(z)^2 - z\tilde{f}_1(z)^2.$$

We will make use of this in our analysis of 2-protected nodes in tries and PATRICIA tries.

2-protected nodes in tries

- Recall: X_n can be described by

$$X_n \stackrel{d}{=} X_{I_n} + X_{n-I_n}^* + T_n.$$

2-protected nodes in tries

- Recall: X_n can be described by

$$X_n \stackrel{d}{=} X_{I_n} + X_{n-I_n}^* + T_n.$$

- By splitting the trie,

$$X_n^{(T)} \stackrel{d}{=} \begin{cases} X_{n-1}^{(T)}, & \text{when } I_n = 1 \text{ or } I_n = n - 1; \\ X_{I_n}^{(T)} + X_{n-I_n}^{(T)*} + 1, & \text{otherwise,} \end{cases}$$

with initial conditions $X_0^{(T)} = X_1^{(T)} = 0$. So,

$$T_n = \begin{cases} 0, & \text{when } I_n = 1 \text{ or } I_n = n - 1; \\ 1, & \text{otherwise.} \end{cases}$$

Mean of 2-protected nodes in tries

- With $\chi_k = \frac{2rk\pi i}{\log p}$ and $\frac{\log p}{\log q} = \frac{r}{l}$ and $(r, l) = 1$, let

$$\mathcal{F}[G](x) := \begin{cases} \frac{1}{h} \sum_{k \in \mathbb{Z} \setminus \{0\}} G(-1 + \chi_k) e^{2k\pi i x}, & \text{if } \frac{\log p}{\log q} \in \mathbb{Q}; \\ 0, & \text{if } \frac{\log p}{\log q} \notin \mathbb{Q}, \end{cases}$$

Mean of 2-protected nodes in tries

- With $\chi_k = \frac{2rk\pi i}{\log p}$ and $\frac{\log p}{\log q} = \frac{r}{l}$ and $(r, l) = 1$, let

$$\mathcal{F}[G](x) := \begin{cases} \frac{1}{h} \sum_{k \in \mathbb{Z} \setminus \{0\}} G(-1 + \chi_k) e^{2k\pi i x}, & \text{if } \frac{\log p}{\log q} \in \mathbb{Q}; \\ 0, & \text{if } \frac{\log p}{\log q} \notin \mathbb{Q}, \end{cases}$$

Theorem

We have

$$\frac{\mathbb{E}(X_n^{(T)})}{n} = \frac{pq + 1 - h}{h} + \mathcal{F}[G_1^{(T)}](r \log_{1/p} n) + o(1),$$

where $G_1^{(T)}(-1 + \chi_k) = \Gamma(-1 + \chi_k) \chi_k (\chi_k pq - pq - 1)$.

Variance of 2-protected nodes in tries

Theorem

We have

$$\frac{\mathbb{V}(X_n^{(T)})}{n} = \frac{G_2^{(T)}(-1)}{h} + \mathcal{F}[G_2^{(T)}](r \log_{1/p} n) + o(1).$$

where $G_2^{(T)}(-1 + \chi_k)$ is

$$\begin{aligned} & \Gamma(3 + \chi_k) (-2^{-3-\chi_k} p^2 q^2) + \Gamma(\chi_k + 2) (2p^2 q(1+p)^{-2-\chi_k} + 2pq^2(1+q)^{-2-\chi_k}) \\ & + \Gamma(\chi_k + 1) (-3pq + 2^{-\chi_k} pq - 2^{-1-\chi_k}) + \Gamma(\chi_k) (1 - 2p(1+p)^{-\chi_k} - 2q(1+q)^{-\chi_k}) \\ & + \Gamma(\chi_k - 1) (1 - 2^{1-\chi_k}) - \frac{1}{h} \sum_{j \in \mathbb{Z}} (\chi_j - 1) G_1^{(T)}(\chi_j - 1) (-1 + \chi_{k-j}) G_1^{(T)}(-1 + \chi_{k-j}) \\ & + 2 \sum_{\ell \geq 1} \frac{(-1)^\ell}{\ell!} \frac{p^{1+\ell} + q^{1+\ell}}{1 - p^{1+\ell} - q^{1+\ell}} K_1(\ell + \chi_k - 1) K_1(-1 - \ell) \Gamma(\ell + \chi_k) \end{aligned}$$

Variance of 2-protected nodes in tries

Theorem (continued)

$$+ 2 \sum_{\ell \geq 2} \frac{(-1)^\ell}{\ell!} \frac{K_1(-\ell)}{1 - p^\ell - q^\ell} \left(pq(p^\ell + q^\ell)(\ell - 1)\Gamma(\chi_k + \ell + 1) + \left(1 - \ell(p^\ell + q^\ell)(2pq + p^{-\ell+2} + q^{-\ell+2})\right)\Gamma(\chi_k + \ell) + (p^\ell + q^\ell)(1 - \ell + p^{-\ell+2}\ell + q^{-\ell+2}\ell)\Gamma(\chi_k + \ell - 1) \right)$$

and $G_2^{(T)}(-1)$ is

$$\frac{2p^2q}{(1+p)^2} + \frac{2pq^2}{(1+q)^2} - 2pq - \frac{p^2q^2}{4} + 2p \log(1+p) + 2q \log(1+q) + \frac{1}{2} + h - 2 \log 2$$
$$+ 2 \sum_{\ell \geq 2} (-1)^\ell \frac{K_1(-\ell)}{1 - p^\ell - q^\ell} \left((p^\ell + q^\ell) \left(\frac{1 - \ell + p^{-\ell+2}\ell + q^{-\ell+2}\ell}{\ell(\ell - 1)} + pq\ell - 3pq - p^{-\ell+2} \right) \right)$$

Variance of 2-protected nodes in tries

Theorem (continued)

$$-q^{-\ell+2}) + \frac{1}{\ell}) + 2 \sum_{\ell \geq 1} \frac{(-1)^\ell}{\ell} \frac{p^{1+\ell} + q^{1+\ell}}{1 - p^{1+\ell} - q^{1+\ell}} K_1(\ell - 1) K_1(-1 - \ell) - \frac{1}{h} (pq + 1 - h)^2$$
$$- \begin{cases} \frac{1}{h \log p} \sum_{j \geq 1} \frac{4rj\pi^2}{\sinh\left(\frac{2rj\pi^2}{\log p}\right)} \left(p^2 q^2 \left(\frac{2rj\pi}{\log p} \right)^2 + (pq + 1)^2 \right), & \text{if } \frac{\log p}{\log q} \in \mathbb{Q}; \\ 0, & \text{if } \frac{\log p}{\log q} \notin \mathbb{Q}. \end{cases}$$

Here, we set

$$K_1(s) := -1 + pqs(s+1) - p^{-s}s - q^{-s}s$$

Variance of 2-protected nodes in tries

Theorem (continued)

$$-q^{-\ell+2}) + \frac{1}{\ell}) + 2 \sum_{\ell \geq 1} \frac{(-1)^\ell}{\ell} \frac{p^{1+\ell} + q^{1+\ell}}{1 - p^{1+\ell} - q^{1+\ell}} K_1(\ell - 1) K_1(-1 - \ell) - \frac{1}{h} (pq + 1 - h)^2$$
$$- \begin{cases} \frac{1}{h \log p} \sum_{j \geq 1} \frac{4rj\pi^2}{\sinh\left(\frac{2rj\pi^2}{\log p}\right)} \left(p^2 q^2 \left(\frac{2rj\pi}{\log p} \right)^2 + (pq + 1)^2 \right), & \text{if } \frac{\log p}{\log q} \in \mathbb{Q}; \\ 0, & \text{if } \frac{\log p}{\log q} \notin \mathbb{Q}. \end{cases}$$

Here, we set

$$K_1(s) := -1 + pqs(s+1) - p^{-s}s - q^{-s}s$$

- Our expression is different from J. Gaither and M. D. Ward.

Variance of 2-protected nodes in tries

Theorem (continued)

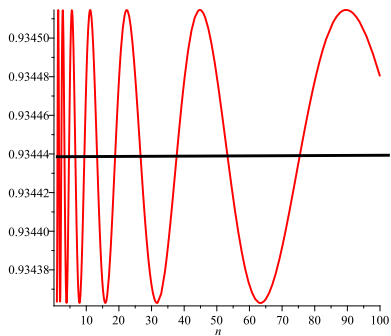
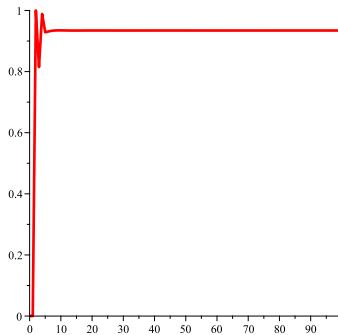
$$-q^{-\ell+2}) + \frac{1}{\ell}) + 2 \sum_{\ell \geq 1} \frac{(-1)^\ell}{\ell} \frac{p^{1+\ell} + q^{1+\ell}}{1 - p^{1+\ell} - q^{1+\ell}} K_1(\ell - 1) K_1(-1 - \ell) - \frac{1}{h} (pq + 1 - h)^2$$
$$- \begin{cases} \frac{1}{h \log p} \sum_{j \geq 1} \frac{4rj\pi^2}{\sinh\left(\frac{2rj\pi^2}{\log p}\right)} \left(p^2 q^2 \left(\frac{2rj\pi}{\log p} \right)^2 + (pq + 1)^2 \right), & \text{if } \frac{\log p}{\log q} \in \mathbb{Q}; \\ 0, & \text{if } \frac{\log p}{\log q} \notin \mathbb{Q}. \end{cases}$$

Here, we set

$$K_1(s) := -1 + pqs(s+1) - p^{-s}s - q^{-s}s$$

- Our expression is different from J. Gaither and M. D. Ward.
- We give an explanation why Abel summability gives the correct result in the expression above.

$$\mathbb{V}(X_n^{(T)})/n$$



Central limit theorem

Theorem (CLT)

We have,

$$\frac{X_n^{(T)} - \mathbb{E}(X_n^{(T)})}{\sqrt{\mathbb{V}(X_n^{(T)})}} \xrightarrow{d} N(0, 1),$$

where \xrightarrow{d} denotes convergence in distribution and $N(0, 1)$ is the standard normal distribution.

2-protected nodes in PATRICIA tries

- By splitting the PATRICIA trie,

$$X_n^{(P)} \stackrel{d}{=} \begin{cases} X_n^{(P)}, & \text{when } I_n = 0 \text{ or } n; \\ X_{n-1}^{(P)}, & \text{when } I_n = 1 \text{ or } n - 1; \\ X_{I_n}^{(P)} + X_{n-I_n}^{(P)*} + 1, & \text{otherwise,} \end{cases}$$

with initial conditions $X_0^{(P)} = X_1^{(P)} = 0$. So,

$$T_n = \begin{cases} 0, & \text{if } I_n \in \{0, 1, n - 1, n\}; \\ 1, & \text{otherwise.} \end{cases}$$

2-protected nodes in PATRICIA tries

- By splitting the PATRICIA trie,

$$X_n^{(P)} \stackrel{d}{=} \begin{cases} X_n^{(P)}, & \text{when } I_n = 0 \text{ or } n; \\ X_{n-1}^{(P)}, & \text{when } I_n = 1 \text{ or } n - 1; \\ X_{I_n}^{(P)} + X_{n-I_n}^{(P)*} + 1, & \text{otherwise,} \end{cases}$$

with initial conditions $X_0^{(P)} = X_1^{(P)} = 0$. So,

$$T_n = \begin{cases} 0, & \text{if } I_n \in \{0, 1, n - 1, n\}; \\ 1, & \text{otherwise.} \end{cases}$$

- By the same analysis as for tries, we get the results of PATRICIA tries.

Mean of 2-protected nodes in PATRICIA tries

Mean of 2-protected nodes in PATRICIA tries

Theorem

We have

$$\frac{\mathbb{E}(X_n^{(P)})}{n} = \frac{pq}{h} + \mathcal{F}[G_1^{(P)}](r \log_{1/p} n) + o(1),$$

where $G_1^{(P)}(-1 + \chi_k) = pq\Gamma(\chi_k + 1)$.

- Here, we have the following relationship between 2-protected nodes in tries and PATRICIA tries via the number of internal nodes:

$$X_n^{(P)} = X_n^{(T)} - N_n + n - 1.$$

Variance of 2-protected nodes in PATRICIA tries

Theorem

We have

$$\frac{\mathbb{V}(X_n^{(P)})}{n} = \frac{G_2^{(P)}(-1)}{h} + \mathcal{F}[G_2^{(P)}](r \log_{1/p} n) + o(1),$$

where $G_2^{(P)}(-1 + \chi_k)$ is

$$\begin{aligned} & \Gamma(3 + \chi_k) (-2^{-3-\chi_k} p^2 q^2) + \Gamma(\chi_k + 2) (2p^2 q(1+p)^{-2-\chi_k} + 2pq^2(1+q)^{-2-\chi_k} - 2^{-1-\chi_k} pq) \\ & + \Gamma(\chi_k + 1) (-3pq - 2^{-1-\chi_k}(1+2pq+p^{1-\chi_k}+q^{1-\chi_k}) + 2p(1+q)(1+p)^{-1-\chi_k} \\ & + 2q(1+p)(1+q)^{-1-\chi_k}) + (\Gamma(\chi_k) + \Gamma(\chi_k - 1))(2(1+p)^{1-\chi_k} + 2(1+q)^{1-\chi_k} - 2^{1-\chi_k} \\ & + (1 - 2^{1-\chi_k})(p^{1-\chi_k} + q^{1-\chi_k}) - 3) + 2 \sum_{\ell \geq 2} \frac{(-1)^\ell}{\ell!} \frac{K_1(-\ell)}{1 - p^\ell - q^\ell} (pq(p^\ell + q^\ell)(\ell - 1) \\ & \Gamma(\chi_k + \ell + 1) + ((p^\ell + q^\ell)(\ell - 2pql - p^{-\ell+2}\ell - q^{-\ell+2}\ell - 1) + 1) \Gamma(\chi_k + \ell) \end{aligned}$$

Variance of 2-protected nodes in PATRICIA tries

Theorem (continued)

$$\begin{aligned} & + (1 - p^\ell - q^\ell)\Gamma(\chi_k + \ell - 1) \Big) - \frac{1}{h} \sum_{j \in \mathbb{Z}} (\chi_j - 1)G_1^{(P)}(\chi_j - 1)(-1 + \chi_{k-j})G_1^{(P)}(-1 + \chi_{k-j}) \\ & + 2 \sum_{\ell \geq 1} \frac{(-1)^\ell}{\ell!} \frac{p^{1+\ell} + q^{1+\ell}}{1 - p^{1+\ell} - q^{1+\ell}} K_1(\ell + \chi_k - 1)K_1(-1 - \ell)\Gamma(\ell + \chi_k), \end{aligned}$$

and $G_2^{(P)}(-1)$ is

$$\begin{aligned} & - \frac{p^2 q^2}{4} - \frac{9pq}{2} - 1 + \frac{2p(1+q)}{1+p} + \frac{2q(1+p)}{1+q} + \frac{2p^2 q}{(1+p)^2} + \frac{2pq^2}{(1+q)^2} \\ & + 2 \sum_{\ell \geq 2} \frac{(-1)^\ell K_1(-\ell)}{1 - p^\ell - q^\ell} \left((p^\ell + q^\ell) \left(1 - p^{-\ell+2} - q^{-\ell+2} - \frac{1}{\ell-1} + pq\ell - 3pq \right) \right) \end{aligned}$$

Variance of 2-protected nodes in PATRICIA tries

Theorem (continued)

$$\begin{aligned} & + \frac{1}{\ell - 1} \Big) + 2 \sum_{\ell \geq 1} \frac{(-1)^\ell}{\ell} \frac{p^{1+\ell} + q^{1+\ell}}{1 - p^{1+\ell} - q^{1+\ell}} K_1(\ell - 1) K_1(-1 - \ell) - \frac{p^2 q^2}{h} \\ & - \begin{cases} \frac{1}{h \log p} \sum_{j \geq 1} \frac{4rj\pi^2}{\sinh\left(\frac{2rj\pi^2}{\log p}\right)} \left(p^2 q^2 \left(\frac{2rj\pi}{\log p} \right)^2 + p^2 q^2 \right), & \text{if } \frac{\log p}{\log q} \in \mathbb{Q}; \\ 0, & \text{if } \frac{\log p}{\log q} \notin \mathbb{Q}. \end{cases} \end{aligned}$$

Here, we set $K_1(s) := (pqs + 1 - p^{-s} - q^{-s})(s + 1)$.

Bivariate central limit theorem

Theorem

We have

$$\frac{\text{Cov}(N_n, X_n^{(T)})}{n} = \frac{H_2(-1)}{h} + \mathcal{F}[H_2](r \log_{1/p} n) + o(1),$$

where

$$H_2(x) = \frac{G_2^{(T)}(x) + G_2^{(N)}(x) - G_2^{(P)}(x)}{2}.$$

Bivariate central limit theorem

Theorem

We have

$$\frac{\text{Cov}(N_n, X_n^{(T)})}{n} = \frac{H_2(-1)}{h} + \mathcal{F}[H_2](r \log_{1/p} n) + o(1),$$

where

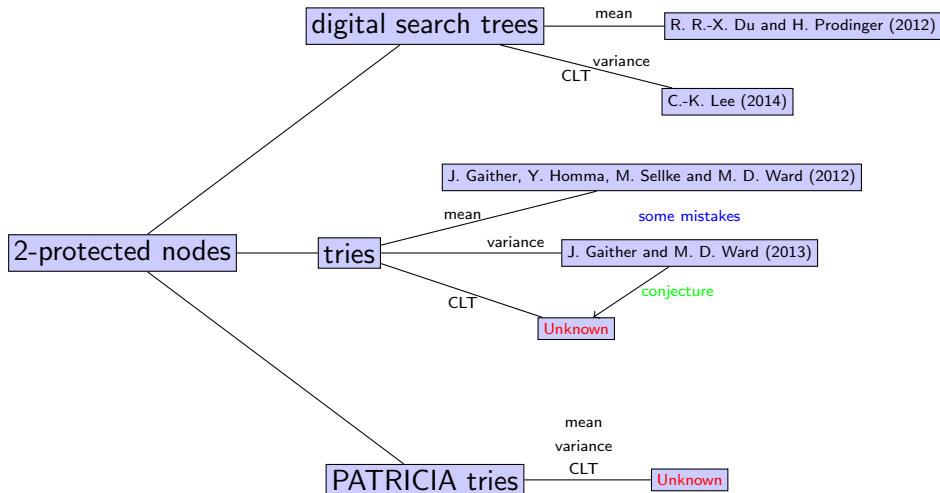
$$H_2(x) = \frac{G_2^{(T)}(x) + G_2^{(N)}(x) - G_2^{(P)}(x)}{2}.$$

Theorem

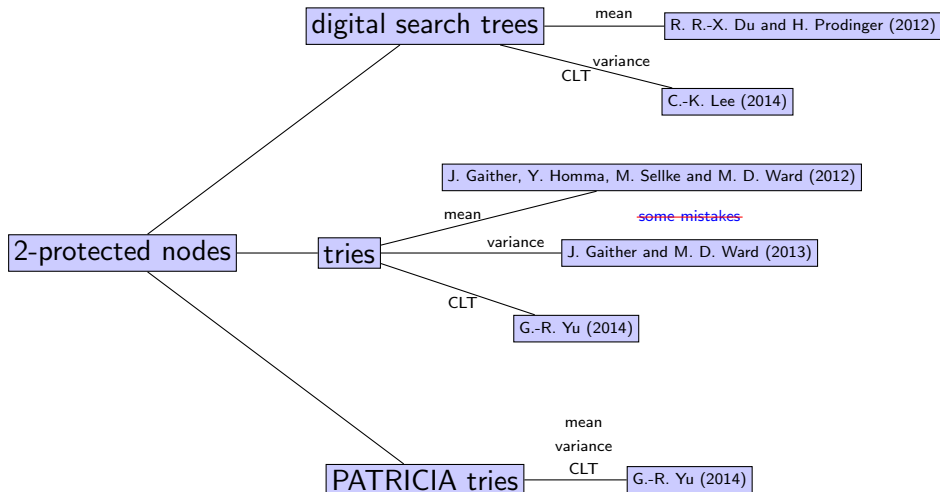
We have,

$$\Sigma_n^{-1/2} \begin{pmatrix} N_n - \mathbb{E}(N_n) \\ X_n^{(T)} - \mathbb{E}(X_n^{(T)}) \end{pmatrix} \xrightarrow{d} N(0, I_2),$$

Number of 2-protected nodes in digital tree



Number of 2-protected nodes in digital tree



Thanks for your attention!